

Heritability and GWAS Studies for Monocyte–Lymphocyte Ratio

Bochao D. Lin,¹ Gonneke Willemsen,¹ Iryna O. Fedko,¹ Rick Jansen,² Brenda Penninx,² E. de Geus,¹ C. Kluft,³ JoukeJan Hottenga,¹ and Dorret I. Boomsma¹

¹Netherlands Twin Register, Department of Biological Psychology, Vrije Universiteit, Amsterdam, the Netherlands

²Department of Psychiatry, VU Medical Center, Amsterdam, the Netherlands

³Good Biomarker Sciences, Leiden, the Netherlands

The monocyte–lymphocyte ratio (MLR) is a useful biomarker for disease development, but little is known about the extent to which genetic and environmental factors influence MLR variation. Here, we study the genetic architecture of MLR and determine the influence of demographic and lifestyle factors on MLR in data from a Dutch non-patient twin-family population. Data were obtained in 9,501 individuals from the Netherlands Twin Register. We used regression analyses to determine the effects of age, sex, smoking, and body mass index (BMI) on MLR and its subcomponents. Data on twins, siblings and parents ($N = 7,513$) were analyzed by genetic structural equation modeling to establish heritability and genome wide single nucleotide polymorphism (SNP) data from a genotyped subsample ($N = 5,892$) and used to estimate heritability explained by SNPs. SNP and phenotype data were also analyzed in a genome-wide association study to identify the genes involved in MLR. Linkage disequilibrium (LD) score regression and expression quantitative trait loci (eQTL) analyses were performed to further explore the genetic findings. Results showed that age, sex, and age \times sex interaction effects were present for MLR and its subcomponents. Variation in MLR was not related to BMI, but smoking was positively associated with MLR. Heritability was estimated at 40% for MLR, 58% for monocyte, and 58% for lymphocyte count. The Genome-wide association study (GWAS) identified a locus on *ITGA4* that was associated with MLR and only marginally significantly associated with monocyte count. For monocyte count, additional genetic variants were identified on *ITPR3*, *LPAP1*, and *IRF8*. For lymphocyte count, GWAS provided no significant findings. Taking all measured SNPs together, their effects accounted for 13% of the heritability of MLR, while all known and identified genetic loci explained 1.3% of variation in MLR. eQTL analyses showed that these genetic variants were unlikely to be eQTLs. In conclusion, variation in MLR level in the general population is heritable and influenced by age, sex, and smoking. We identified gene variants in the *ITGA4* gene associated with variation in MLR. The significant SNP-heritability indicates that more genetic variants are likely to be involved.

■ **Keywords:** heritability, MLR, GWAS, age and sex differences, lifestyle

As early as the 1960s, the relevance of the monocyte–lymphocyte ratio (MLR) for disease prediction was pointed out. The first studies focused on infectious diseases, suggesting MLR reflects the balance between effector and host (Hanifin & Cline, 1970), and found MLR to predict the development and progress of tuberculosis (Iqbal et al., 2014). In later studies, the association between MLR and other diseases was studied and the MLR emerged as a predictor for cancer (Nishijima et al., 2015). While MLR is examined in relation to disease, studies examining the causes of variance in MLR in the general population are lacking. To fully understand the role of MLR in disease, it is, however, necessary to understand the factors underlying variation in MLR in the general non-patient population.

We recently showed that heritability plays a role in individual differences in two other lymphoid ratios, the neutrophil–lymphocyte ratio (NLR) and the platelet–lymphocyte ratio (PLR). For NLR, the heritability was moderate (35%), but for PLR, heritability was high (64%), with evidence of the presence of non-additive genetic effects

RECEIVED 23 December 2016; ACCEPTED 10 January 2017. First published online 14 February 2017.

ADDRESS FOR CORRESPONDENCE: D. I. Boomsma, Department of Biological Psychology, Vrije Universiteit, Van der Boerhorststraat 1, 1081 BT, Amsterdam, the Netherlands. E-mail: di.boomsma@vu.nl

(Lin et al., 2016). The first genome-wide association study (GWAS) for these two ratios identified a genome-wide locus on the *HBS1L-MYB* intergenic region associated with PLR (Lin et al., 2016), which has been associated with multiple blood parameters, including platelet count (Lin et al., 2016).

Although the heritability of MLR has not been studied, genetic studies have been conducted for its subcomponents, the monocyte and lymphocyte levels, showing heritability estimates of 56–73% for monocyte levels and 35–66% for lymphocyte levels (Evans et al., 1999; Hall et al., 2000). In addition, linkage and GWAS have pointed to the genetic variants partly responsible for the individual variation in monocyte and lymphocyte levels. GWASs have identified eight loci associated with monocyte levels: *ITGA4* at 2q31.3, *HLA-DRB1* at 6p21.32, *CCBP2* at 3p22.1, *RPN1* at 3q21.3, *LPAR1* at 9q31.3, intergenic regions at 8q24 and 3q21, and *IRF8* at 16q24.1 (Crosslin et al., 2013; Keller et al., 2014; Nalls et al., 2011; Reiner et al., 2011). GWAS also identified two loci for lymphocyte levels: 6p21 (*EPS15L1* gene) and 19p13 (Crosslin et al., 2013; Nalls et al., 2011). Taken together, genetic factors are likely to play a role in normal variation in MLR, but the nature of the involvement remains to be determined.

Environmental and lifestyle factors may also influence MLR levels. Here, too, studies on MLR itself are lacking, but our own study on PLR and NLR (Lin et al., 2016) showed that lifestyles influence PLR and NLR levels, and agreed with other studies (Azab et al., 2014; Li et al., 2015) that smoking and BMI may also affect these parameters. In addition, its subcomponents, monocyte and lymphocyte levels, have been found to be influenced by BMI (Tenorio et al., 2014; Zaldivar et al., 2006) and smoking (Schwartz & Weiss, 1994), although these effects are not found in all studies (Al-Sufyani & Mahassni, 2011; Perez-de-Heredia et al., 2015).

In this article, we examined several potential causes of variance in MLR in the general population. First, age, sex, smoking, and BMI effects on MLR were investigated. Then, we studied genetic influences on MLR to provide more insight into its genetic architecture in a healthy population. We started by estimating the heritability of MLR and its subcomponents by extended twin family modeling. Next, we used GWAS to identify genetic variants associated with MLR variation and genome-wide complex trait analysis (GCTA) to determine the percentage of variance of MLR that is explained by significant versus all measured genetic variants (single nucleotide polymorphisms; SNPs). Subsequently, we performed an expression quantitative trait loci (eQTL) analysis of all the top SNPs, which were significantly associated with MLR. We repeated the GWAS, GCTA, and eQTL analyses for monocyte levels, while referring to lymphocyte levels in the results presented previously (Lin et al., 2016). Finally, linkage disequilibrium (LD) score regression was performed using the summary statistics of the GWAS results to determine the polygenetic effects and genetic correlations between MLR and subcomponents.

Methods

Participants and Phenotypes

All participants were adults registered with the Netherlands Twin Register (NTR), who took part in a longitudinal study on health and lifestyle in twins and their family members (Willemsen et al., 2013). Data were obtained as part of NTR biobanking projects conducted in 2004–2008 (Willemsen et al., 2010, 2013; Sirota et al., 2015). After removing outliers (i.e., absolute values exceeding mean ± 5 SD), data on monocyte and lymphocyte count were available for 9,501 participants clustered in 3,412 families. During the interview conducted at the time of the home visit, height and weight were obtained. BMI was calculated as weight (kg) divided by height squared (m²). Participants reported whether they currently smoked or had smoked. If so, they were asked for the number of cigarettes smoked per day and how long they (had) smoked. Participants were divided into five categories: non-smoker, ex-smoker, light smoker (currently smoking less than 10 cigarettes a day), average smoker (currently smoking 10–19 cigarettes a day), and heavy smoker (currently smoking 20 or more cigarettes a day). Participants were asked to indicate when they were last ill and the nature of the illness. In the case of medication use, the dosage, brand, and name were recorded. In addition, participants indicated on the presence and nature of any chronic disease. The following exclusion criteria were used to identify individuals who may have had a compromised immune system at the time of blood sampling: (1) illness reported in the week prior to sampling ($N = 552$); (2) C-reactive protein (CRP) ≥ 15 ($N = 307$); (3) basophile count $> 0.02 \times 10^9/L$ ($N = 154$); (4) report of blood-related disease or cancer ($N = 84$); and (5) use of anti-inflammatory medication ($N = 423$), glucocorticoids ($N = 143$), or iron supplements ($N = 29$). Participants meeting one or more of these criteria were labeled as unhealthy ($N = 1,362$), leaving 8,139 individuals from 3,280 families as the population that we will here refer to as the healthy population. Genetic twin-family modeling was conducted using data from twin families limited to at most one twin pair per family and at most two brothers and two sisters and father and mother. This resulted in a sample of 7,513 participants from 3,252 families, including 240 monozygotic male (MZM), 98 dizygotic same-sex male (DZM), 536 monozygotic female (MZF), 219 dizygotic same-sex female (DZF), and 222 dizygotic opposite-sex (DOS) twin pairs. The study protocol was approved by the Medical Ethics Committee of the VU University Medical Center Amsterdam and all participants provided informed consent.

Procedure

Participants were visited at home, or when preferred, at work, to obtain blood samples and conduct a brief health-related interview. Visits took place in the morning between 7 am and 10 am and women were seen, when possible,

between the 2nd and the 4th day of the menstrual cycle or, if on hormonal birth control, were visited in their pill-free week. Participants were asked to fast from the evening before and to refrain from smoking or physical exercise 1 hour before blood sampling (for more details see Willemssen et al., 2010). Peripheral blood was collected in anticoagulant vacuum tubes, which were inverted 8–10 times immediately after the blood draw. All samples were transported to the laboratory facility in Leiden, the Netherlands, within 3–6 hours after blood sampling. The blood samples were then directly used or stored to measure parameters of interest or extract DNA or RNA at a later moment (see Lin et al., 2016; Willemssen et al., 2010). The hematological profile was obtained from EDTA blood samples with a Coulter system (Coulter Corporation Miami, USA). This profile consisted of total white blood cell count, percentages and numbers of neutrophils, lymphocytes, monocytes, eosinophils, and basophils, and indicators of red blood cell types and platelets. We calculated MLR as the absolute monocyte count ($10^9/L$) divided by the absolute lymphocyte count ($10^9/L$). CRP was determined from a heparin plasma sample using the 1,000 CRP assay (Diagnostic Product Corporation).

Genotype Data

For DNA isolation, we used the GENTRA Puregene DNA isolation kit on frozen whole blood samples, which were collected in EDTA tubes. All procedures were performed according to the manufacturer's protocols (Boomsma et al., 2008). Genotyping was done on multiple platforms, including a number of partly overlapping subsets of participants. The following platforms were used: Affymetrix Perlegen 5.0, Illumina 370, Illumina 660, Illumina Omni Express 1 M, and Affymetrix 6.0 (for details see (Lin et al., 2016)). The individual SNP markers were lifted over to Build 37 (HG19) of the Human Reference Genome using the LiftOver tool (<http://genome.sph.umich.edu/wiki/LiftOver>). Genotype calls were made with platform specific software (BIRD-SUITE APT-Genotyper Beadstudio) for each specific array. Phasing of all samples and imputing cross-missing platform SNPs was done with MACH 1 (Li et al., 2010). The phased data were then imputed with MINIMAC (Howie et al., 2012) in batches of around 500 individuals for the autosomal genome using the above 1000G Phase I integrated reference panel for 561 chromosome chunks obtained by the CHUNKCHROMOSOME program (Liu et al., 2013). SNPs were removed if the Mendelian error rate was > 0.02 , if the imputed allele frequency differed more than 0.15 from the 1000G reference allele frequency, and if $MAF < 0.01$ and if $R^2 < 0.80$. Hardy–Weinberg Equilibrium was calculated on the genotype probability counts for the full sample and SNPs were removed, if the p value $< .00001$. After imputation, MZ twins were reduplicated back into the data. This left 6,010,458 SNPs in the GWAS analyses.

As several different platforms were used, additional SNP quality control (QC) included an evaluation of the SNP

platform effects, and SNPs showing platform effects were removed. This was done by defining individuals on a specific platform as cases and the remaining individuals as controls. Allelic association was then calculated and SNPs were removed if the specific platform allele frequencies were significantly different from the remaining platforms with p value $< .00001$. In total, 5,987,253 SNPs survived this QC and these SNPs were then used to build the genetic relationship matrix (GRM) for all individuals. The selected SNPs were transformed to best guess Plink binary format and subsets were made for each of the 22 chromosomes. The GRMs for all NTR samples were then calculated using GCTA (Yang et al., 2011). We generated 24 GRMs in total. A first autosomal GRM reflects an identity-by-state (IBS) matrix for all individuals. This GRM is determined from all autosomal SNPs and is used to estimate the SNP heritability (h_g^2). A second autosomal GRM represents closely related individuals (identity-by-state [IBS] > 0.05), so any remaining pairwise relationship estimates smaller than 0.05 were set to zero in this matrix. This matrix is used as second covariate matrix in the GWAS and GCTA studies to account for the family structure of individuals and to estimate the narrow-sense heritability (h^2) of applying an additive model. Finally, 22 GRMs were created that included all autosomal SNPs except for those on the one chromosome that is tested in the GWAS (the leave one chromosome out or LOCO strategy). These matrices were used in the GWAS as a covariate matrix to remove artificial inflation due to all kinds of subsample stratification.

Statistical Analyses

First, using age- and sex-corrected values, we tested for differences in MRL, monocyte, and lymphocyte levels between the healthy and unhealthy population using a t -test. Next, within the healthy population, we explored the age and sex effects by linear regression. To detect the influence of lifestyle on variation in the MLR level, we included BMI and smoking behavior in a regression analysis conducted separately by sex, taking age into account. Analyses were conducted in STATA (Stata Corp., 2013) using the cluster option to correct for the family structure within the data. Using genetic structural equation modeling in OpenMx (Boker et al., 2011), the heritability of MLR, monocyte count, and lymphocyte count was estimated in the healthy population. MLR, monocyte count, lymphocyte count, and age were standardized using z scores. Parameters were estimated by maximum likelihood. We summarized the family resemblance with respect to MLR by means of correlations corrected for age, sex, and age \times sex effects. Then, we fitted a series of genetic models. The total phenotypic variance was decomposed into four sources of variation: additive genetic (A), non-additive genetic (D), common environmental (C), and unique environmental (E) variation. The common environmental variance reflects the variance shared between siblings and twins (V_C). The resemblance among

family members was modeled as a function of A, D, and C. We allowed for a correlation in phenotype between spouses (μ). In fitting the genetic models, we included as covariates age, sex, sex \times age. We fitted the full model as described and tested the presence of assortative mating (i.e., the correlation between phenotypes of spouses) and the presence of shared environment and non-additive influences. The nested submodels were compared to the full model by log likelihood ratio test ($-2LL$) using a significance level of 0.05.

GWAS

We performed two GWASs: one for MLR level and one for monocyte count, using the quality controlled imputed SNPs, including age, sex, three Dutch principal components (PCs) generated with the Eigensoft software and genotype platform as covariates ($N = 5,892$; see Methods and Abdellaoui et al., 2013). As we had already conducted a GWAS for lymphocyte count, using a largely overlapping sample ($N = 5,901$, overlap of 5,890 individuals) we did not rerun this analysis but instead refer here to the results published in Lin et al. (2016). Analyses were performed with the GCTA software running a mixed linear model association model to account for relatedness (Tucker et al., 2015). To avoid inflated test statistics in datasets with related individuals and other remaining cryptic stratification, we used two covariate GRM matrices: the matrix for all individuals, excluding the chromosome under analysis (LOCO analysis) and the matrix only focusing related individuals with $IBS > 0.05$ (Tucker et al., 2015). For the GWAS, we assumed the statistically significant threshold as a p value less than 5×10^{-8} (Pe'er et al., 2008), and we refer to as marginally significant when p values exceed this threshold but remain below 10^{-4} .

eQTL Analysis

To detect possible causal effects for significant genetic variants, we conducted eQTL analyses, in which eQTL effects were detected with a linear model approach using Matrix-eQTL (Shabaln, 2012). The analysis specified expression level as the dependent variable and SNP genotype values as the independent variable. The eQTL dataset is described in detail elsewhere (Jansen et al., in press; Wright et al., 2014). eQTL effects were defined as cis when the distance between probe set–SNP pairs was smaller than 1M base pairs (Mb), and as trans when the SNP and the probe set were separated by more than 1 Mb on the genome according to hg19.

GCTA and Linkage Disequilibrium Score Regression

We performed GCTA analyses to estimate narrow-sense heritability, the fraction of genetic variance explained by the significant SNPs detected in the GWAS and the fraction of genetic variance explained by the known significant SNPs from the published literature. These analyses were done for MLR level, monocyte count, and lymphocyte count. A restricted maximum likelihood analysis procedure was used under a linear design (Yang et al., 2011). Sex, age, genotype

TABLE 1

Average Levels (SD) for MLR and Its Subcomponents in the Healthy and Unhealthy Population for Men and Women

	Healthy population		Unhealthy population	
	Men	Women	Men	Women
<i>N</i>	3,074	5,065	444	918
Age	44.12 (15.88)	42.98 (14.48)	46.75 (17.37)	43.75 (15.48)
MLR	0.29 (0.09)	0.24 (0.08)	0.32 (0.14)	0.26 (0.12)
Monocyte	0.58 (0.17)	0.51 (0.16)	0.61 (0.19)	0.54 (0.20)
Lymphocyte	2.17 (0.64)	2.27 (0.71)	2.07 (0.75)	2.22 (0.75)

platform, and three Dutch PCs were included as covariates. We used two covariance matrixes to estimate narrow sense heritability (h^2), and GWAS and known loci heritability (Zaitlen et al., 2013). The first GRM is the full autosomal GRM as described previously. The second GRM is the closely related ($IBS > 0.05$) matrix. Pearson correlations between the phenotypes of interest were calculated in R (R Core Team., 2014). Whether polygenetic effects (Bulik-Sullivan et al., 2015) influenced MLR and its compositions was explored by LD score regression. The SNP heritability (Finucane et al., 2015) of MLR, monocyte count, lymphocyte count, and genetic correlations (Bulik-Sullivan et al., 2015) among the phenotypes were determined by LD score regression on our computed GWAS summary statistics. The genetic correlation of two traits can be calculated by the slope from the LD regression on the product of effect sizes (z score) for two phenotypes of interest. In order to do this, we used the HapMap3 LD scores (N SNPs = 1,293,150) computed for each SNP based on the LD observed in European ancestry individuals from the 1,000 Genomes project (accessible online at <http://github.com/bulik/ldsc>). QC for genetic data is the default setting in the program.

Results

Health Status, Sex, Age, and Lifestyle

Table 1 gives the descriptive statistics for MLR and its subcomponents, the monocyte and lymphocyte count, for the healthy and unhealthy parts of the population. The comparison of the healthy and unhealthy population (see Methods for definition), taking sex and age into account as well as family structure, showed, as expected, that individuals in the unhealthy population had on average a higher MLR ratio, $t(9,499) = -7.95$, $p < .001$, and monocyte count, $t(9,499) = -5.06$, $p < .001$, and a lower lymphocyte count, $t(9,499) = -2.57$, $p = .01$). We continued our investigation in the healthy population, examining the influence of age and sex. Men had higher MLR levels than women, $b = -0.0176$, $p < .01$, and MLR increased with age in both men and women, $b = 0.0013$, $p < .001$. There was also evidence for an age \times sex interaction: the age effects were alleviated in the women. With respect to the subcomponents, monocyte, and lymphocyte levels were higher in men than

TABLE 2
Results of the Linear Regression Analyses (Regression Coefficients) for MLR and Its Subcomponents for Men and Women

Dependent variable	Independent variable	Men		Women	
		Model 1	Model 2	Model 1	Model 2
MLR	Age	0.0013***	0.0016**	0.0008***	0.0014***
	BMI	–0.0041	–0.0011	–0.0034	0.0075
	Smoking	–0.0029*	0.0016	–0.0044***	0.0046
	Age × BMI		0.0007		–0.0002*
	Age × smoking		–0.0001		2.8E-6
Monocyte count	Age	0.00016	0.0005	–0.0001	–0.0011
	BMI	0.0193***	0.0304*	0.0151***	–0.0003
	Smoking	0.0343***	0.0245**	0.0297***	0.00262***
	Age × BMI		–0.0002		0.0004
	Age × smoking		0.0002		0.0009
Lymphocyte count	Age	–0.0078***	–0.0086**	–0.0078***	–0.0154***
	BMI	0.0947***	0.0972*	0.0902***	–0.0487
	Smoking	0.1512***	0.1163***	0.1765***	0.1836***
	Age × BMI		–0.0003		0.0031**
	Age × smoking		0.0009		–0.0001

Note: Estimates in bold are significant (* $p < .05$, ** $p < 0.01$, *** $p < .001$). MLR = monocyte count ratio; BMI = body mass index; Age × BMI = interaction of age and BMI; Age × smoking = interaction of age and smoking; Model 1: blood variable = $b_0 + b_1 \times \text{age} + b_2 \times \text{BMI} + b_3 \times \text{smoking}$; Model 2: blood variable = $b_0 + b_1 \times \text{age} + b_2 \times \text{BMI} + b_3 \times \text{smoking} + b_4 \times \text{age} \times \text{BMI} + b_5 \times \text{age} \times \text{smoking}$.

in women and increased with age. These age effects were similar in men and women. To test the effects of BMI and smoking, we included this variable in a regression analysis conducted separately by sex and taking age into account. The results, shown in Table 2 (model 1), indicate that smoking is related to a decrease in MLR level in both men and women. BMI was not associated with MLR in either sex. However, an age × BMI interaction was seen for MLR in women (model 2): the age effects were alleviated by an increased BMI level. The BMI and smoking effects were also examined in the MLR subcomponents: Higher BMI and being a smoker were related to higher monocyte and lymphocyte levels. For lymphocyte count in women, there was evidence for an age × BMI interaction, again indicating a reduction of the BMI effect at an older age.

Heritability

The known genetic relations among mono- and di-zygotic twins and their family members were used to model familial resemblance in MLR, and monocyte and lymphocyte count as a function of genetic and environmental parameters. These models included sex, age, and sex × age effects as fixed effects. Table 3 contains the familial correlations as obtained for MLR, monocyte, and lymphocyte count. For MLR, twin pair correlations did not depend on sex, and the correlations did not differ across DZ twin and sibling relations. The correlations in MZ males and MZ female twin pairs were equal as were the other male and female first-degree relative correlations. The resulting MZ correlation was 0.43 (CI is 0.33–0.46) and the DZ correlation was 0.22 (0.14–0.24), with spousal correlations significant at 0.104 (0.002–0.135). The pattern of twin correlations showed no evidence for non-additive or common environmental effects. This was confirmed by model fitting in which the heritability of MLR was estimated at 40% (0.34–0.43).

We also conducted these series of genetic modeling analyses for monocyte and lymphocyte count. For monocyte count, there were no significant spousal correlations and the MZ correlation was 0.58 (0.54–0.62) while the DZ correlation was 0.27 (0.21–0.31). In line with the pattern of the correlations, genetic modeling estimated the broad sense heritability at 58%, with non-additive effects accounting for 12% and no evidence for the influence of common environmental factors. For lymphocyte count, we estimated the heritability in the current set ($N = 5,892$, with >99% overlap with the set described in Lin et al., 2016) and, as to be expected, results were similar to those published in Lin et al. (2016) with a broad sense heritability at 58% and non-additive effects accounting for 22%.

GWAS

Figures 1, 2 and 3 show the QQ and Manhattan plots from the GWAS for MLR, monocytes and lymphocytes. After adjusting for age, sex, genotype platform, and PCs, and using the LOCO and family-based GRM correction, the GWAS λ s were 0.9965 for MLR and 1.0166 for monocyte count.

For MLR, associations were found with 11 SNPs situated on the *ITGA4* (VLA-4 subunit alpha) gene on chromosome 2q31 (Figure 1 and Table 4). The top SNP rs3755021 T allele was linked to a decrease in MLR level ($\beta = -0.012$, $p = 2.21E-8$). This SNP was not associated with lymphocyte count, but in our study was marginally significantly associated with monocyte count ($\beta = -0.018$, $p = 6.34E-6$), and has also been associated with monocyte count in a linkage study (Maugeri et al., 2011) and two previous GWA studies (Crosslin et al., 2013; Nalls et al., 2011). The G allele of rs6740847 in this region has been linked to decreased *ITGA4* expression levels in blood, which increases the number of circulating monocytes and may indicate this is a causal gene (Maugeri et al., 2011).

TABLE 3
Familial Correlations (Confidence Interval) for MLR Monocyte and Lymphocyte Count Within the Healthy Population

Pairs	MLR		Monocyte		Lymphocyte	
	R	95% CI	R	95% CI	R	95% CI
MZ twins	0.431	0.330–0.463	0.583	0.539–0.622	0.582	0.537–0.621
MZ male	0.340	0.329–0.344	0.515	0.432–0.585	0.576	0.487–0.646
MZ female	0.489	0.478–0.492	0.620	0.568–0.663	0.584	0.531–0.630
Male first-degree relatives	0.182	0.082–0.205	0.247	0.183–0.308	0.234	0.156–0.307
DZ male	0.244	0.121–0.248	0.170	–0.034–0.354	0.390	0.153–0.551
Brother–male twin	0.128	0.116–0.132	0.182	0.054–0.301	0.181	0.007–0.336
Brother–brother	0.242	0.231–0.246	0.346	0.082–0.545	0.341	0.007–0.580
Father–son	0.186	0.174–0.190	0.269	0.194–0.339	0.219	0.122–0.307
Female first-degree relatives	0.224	0.139–0.239	0.228	0.186–0.268	0.225	0.183–0.266
DZ female	0.279	0.267–0.282	0.385	0.154–0.279	0.286	0.181–0.382
Sister–female twin	0.169	0.157–0.173	0.285	0.156–0.390	0.175	0.008–0.265
Sister–sister	0.241	0.230–0.247	0.315	0.214–0.406	0.151	0.055–0.243
Mother–daughter	0.228	0.216–0.232	0.247	0.183–0.307	0.210	0.157–0.260
Female–male first degree relatives	0.225	0.152–0.235	0.285	0.237–0.331	0.203	0.162–0.244
DZ opposite sex	0.098	0.086–0.103	0.181	0.054–0.301	0.216	0.086–0.333
Brother–female twin	0.232	0.220–0.236	0.335	0.026–0.443	0.182	0.007–0.336
Sister–male twin	0.193	0.181–0.197	0.189	0.060–0.307	0.172	0.023–0.307
Sister–brother	0.200	0.188–0.204	0.227	0.080–0.352	0.312	0.178–0.419
Mother–son	0.217	0.205–0.221	0.240	0.168–0.308	0.198	0.115–0.276
Father–daughter	0.262	0.250–0.265	0.218	0.154–0.278	0.245	0.183–0.260
Parents (father–mother)	0.104	0.002–0.135	0.061	–0.013–0.135	0.166	0.089–0.241
Narrow-sense heritability(V_A)	0.397	0.341–0.429	0.468	0.404–0.530	0.353	0.294–0.402
Broad-sense heritability(V_A+V_D)			0.583	0.446–0.720	0.576	0.461–0.653

Note: Correlations in bold type were obtained from submodels in which all matching correlations of the tested subgroup of family relations were constrained to be equal.

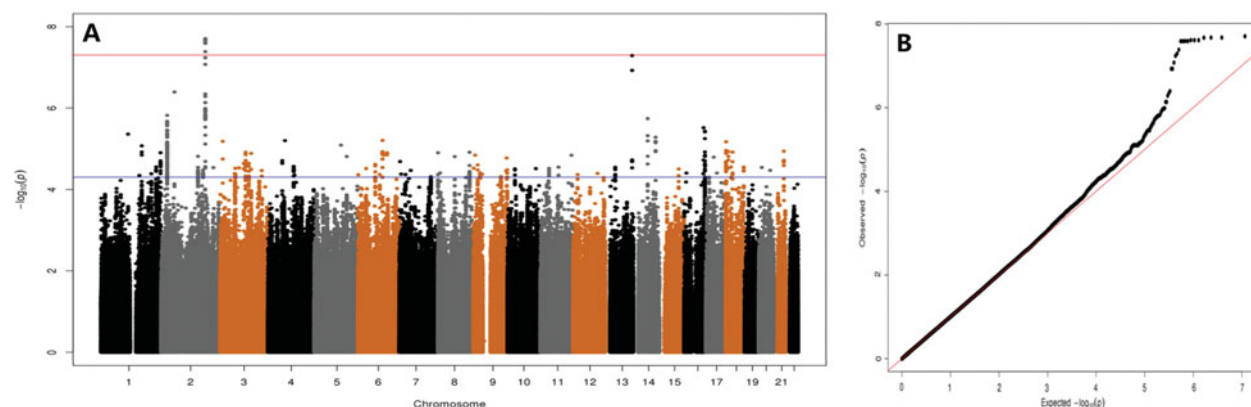


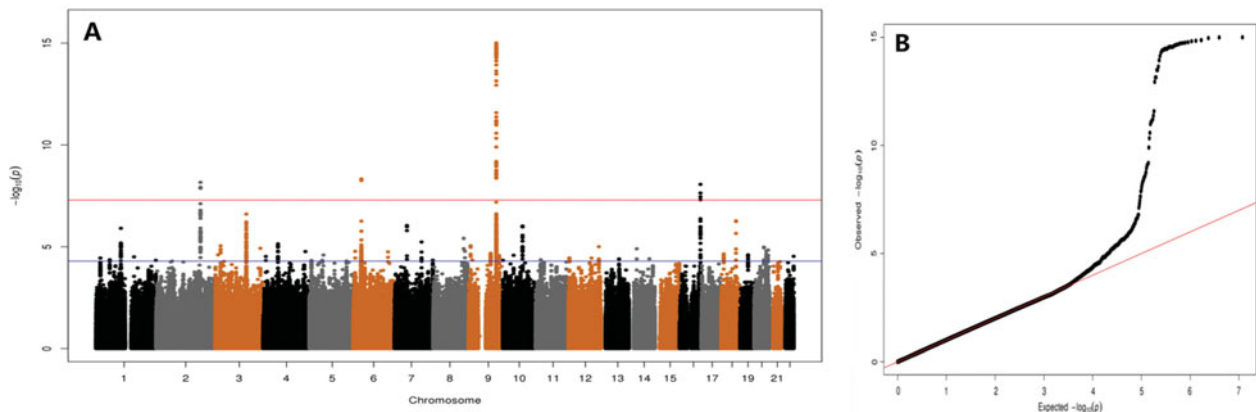
FIGURE 1
(Colour online) Manhattan and QQ plot for MLR level with SNPs having a minor allele frequency above 0.01 ($\lambda = 0.996503$).

For monocyte count, the four top hits were rs13029501 at *ITGA4*, rs55929401 located at a region nearby *LPAR1* at 9q31.3, rs391855 at *IRF8*, and rs9469532 at 6p21. The most significant locus rs13029501 at 9q31 has been previously associated with monocyte count in European and Japanese populations (Ferreira et al., 2009; Kamatani et al., 2010; Maugeri et al., 2011; Nalls et al., 2011). This SNP is 14.6kb upstream of the SNP most significant associated MLR level (rs3755021, the pair wise LD is 0.22). It is located in a region 163kb downstream of lysophosphatidic acid receptor 1 gene (*LPAR1*, also known as *EDG2*) and increases *LPAR1* expression, which is linked to an increased number of monocytes (Maugeri et al., 2011). As indicated previously, genetic variants nearby the *ITGA4* region are

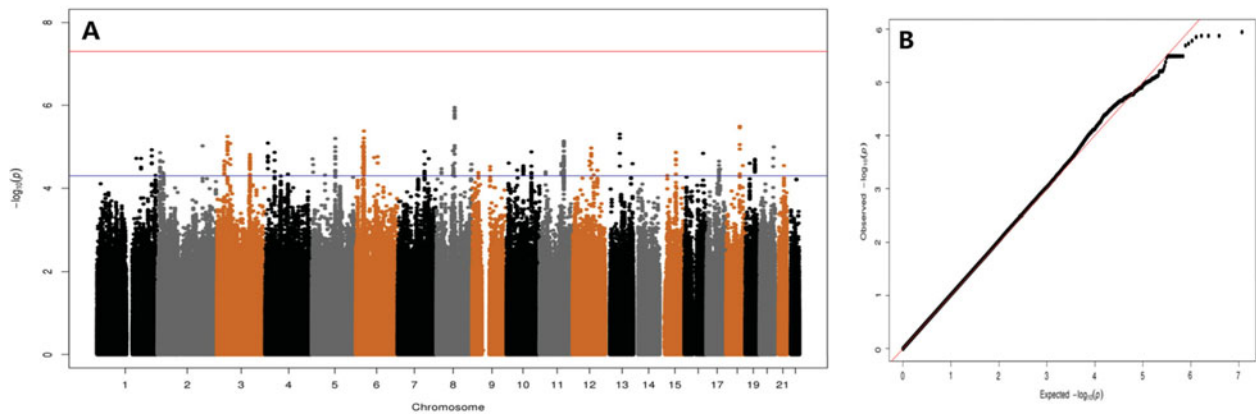
involved in the down regulation of *ITGA4* expression, which increases the number of monocytes circulating in peripheral blood. The *IRF8* gene has also been associated before with monocyte count and has been identified as multiple sclerosis susceptibility loci (De Jager et al., 2009). Animal model studies showed that *IRF8* as a transcription factor plays an essential role in the regulation of lineage commitment during monocyte differentiation (Kurotaki et al., 2013; Terry et al., 2015; Yanez et al., 2015). The top SNP at 6p21 rs9469532 is an intergenic genetic variant nearby *ITPR3*, *LOC101929188*, and *LOC105375023*. The *HLA-DRB1* region 1,043kb upstream of this SNP has previously been associated with monocyte count (Okada et al., 2011).

TABLE 4
Significant SNP Associations for MLR

rs number	Chr	BP	MAF	β (MLR)	SE (MLR)	p (MLR)	p (monocyte count)	p (lymphocyte count)
rs3755021	2	182349409	0.186439	-0.0119024	0.00212769	2.21E-08	6.34E-6	.067
rs17224699	2	182348554	0.186524	-0.0118630	0.00212596	2.40E-08	6.55E-6	.069
rs17290693	2	182350489	0.186354	-0.0118593	0.00212614	2.43E-08	5.63E-6	.074
rs79965377	2	182351172	0.186354	-0.0118593	0.00212614	2.43E-08	5.63E-6	.074
rs17290351	2	182345875	0.187712	-0.0117619	0.00211863	2.83E-08	8.91E-6	.067
rs17224524	2	182346146	0.187712	-0.0117619	0.00211863	2.83E-08	8.91E-6	.067
rs2305588	2	182347072	0.187712	-0.0117619	0.00211863	2.83E-08	8.91E-6	.067
NA	2	182348141	0.187797	-0.0117540	0.00212030	2.96E-08	1.05E-5	.064
rs12479308	2	182348342	0.187797	-0.0117540	0.00212030	2.96E-08	1.06E-5	.064
rs2305590	2	182346544	0.187627	-0.0117502	0.00212048	3.00E-08	9.12E-6	.068
rs2305589	2	182346937	0.187627	-0.0117502	0.00212048	3.00E-08	9.12E-6	.068

**FIGURE 2**

(Colour online) Manhattan and QQ plot for monocyte count with SNPs having a minor allele frequency above 0.01 ($\lambda = 1.0166495$).

**FIGURE 3**

(Colour online) Manhattan and QQ plot for lymphocyte count with SNPs having a minor allele frequency above 0.01 ($\lambda = 1.022341$).

As published in Lin et al. (2016), there were no significant hits when conducting the GWAS for lymphocyte count. However, it is of interest to note that the locus on chromosome 6p21, which was associated with the monocyte level, was also marginally associated with MLR (for

rs9469532, $\beta = -0.069$, $p = 7.69E-5$) and lymphocyte count (for rs114641912, $\beta = -0.059$, $p = 6.19E-6$). This region harbors candidate genes like *ITPR3* (Oishi et al., 2008) and *HLA-DRB1* (Farragher et al., 2008), which have been previously implicated in immunological diseases. In addition,

TABLE 5
Known Blood Cell Count Loci and their Significance in Our GWAS

Phenotype	Gene	SNP	Chr	Location	p	p (mlr)	p (mono)	p (lymp)	
Monocyte	ITGA4	rs2124440	2	182328214	4.63E-10	2.21E-6	1.53E-7	0.838	
	ITGA4	rs1449263	2	182319301	6.71E-14	2.40E-6	2.84E-7	0.776	
	RPN1	rs2712381	3	128338600	1.94E-10	0.005	4.01E-4	0.620	
	c3orf27	rs9880192	3	128297569	1.35E-8	5.97E-7	1.39E-7	0.430	
	CCBP2	rs2228467	3	42906116	1.57E-10	0.769	0.770	0.804	
	CCBP2	rs2228468	3	42907112	5.15E-14	0.128	0.146	0.588	
	Intergenic	rs2047076	5	76058509	1.64E-8	0.492	0.296	0.956	
	HLA-DRB1	rs3095254	6	31221668	8.27E-9	0.910	0.198	0.276	
	Intergenic	rs1991866	8	130624105	4.58E-11	2.44E-6	2.08E-4	0.171	
	EDG2	rs10980800	9	113915905	1.1E-14	2.23E-5	1.28E-14	0.011	
	PTGR1	rs2273788	9	114348617	4.50E-10	0.579	0.024	0.159	
	LPAR1	rs7023923	9	113925534	8.9E-6	4.26E-4	3.81E-12	0.001	
	IRF8	rs424971	16	85946450	3.16E-10	2.1E-4	8.16E-6	0.736	
	Lymphocyte	LOC101929772	rs2524079	6	31242174	1.85E-8	0.0870	0.056	3.02E-4
		EPS15L1	rs11878602	19	16555153	3.42E-9	0.0735	0.857	0.107

Note: Chr = chromosome; p = p values from previous studies; p (mlr) = p values for MLR in current study; p (mono) = p values for monocyte count in current study; p (lymp) = p values for lymphocyte count in current study.

TABLE 6
Overview of eQTL Results: The Association Between Genetic Variants of Interest (β) with Gene Expression Level, Uncorrected for Blood Composition

GWA SNP of interest	Top SNP in eQTL analysis	Gene	LD r^2	β	FDR
rs3755021	rs2305591	ITGA4	0.32	0.191	1.4e-05
rs13029501	rs2305591	ITGA4	0.03	0.203	1.34e-05
rs13029501	rs16867443	CERKL	0.43	0.132	1.34e-05
rs9469532	rs115378869	HLA-DPB1	0.04	0.106	3.07e-04
rs391855	rs1568391	IRF8	0.45	0.13	1.34e-05
rs55929401	rs7023923	LPAR1	0.24	0.42	1.34e-05

Note: LD r^2 : LD between GWAS SNP and top SNP in eQTL analysis. β = eQTL β of GWA SNP, FDR = eQTL FDR GWA SNP.

other loci with ‘potential association peaks’, meaning p values are low but do not reach the required significance level, have been found to be associated with immune disease such as *ERAP1* at 5q15 (Alvarez-Navarro et al., 2015) and *CNTN5* at 11q22 (Thomas et al., 2014).

Table 5 shows the loci for monocyte and lymphocyte count found in previous studies and their significance levels for MLR and its subcomponents in the current study. For some loci, p values were low, indicating a ‘potential’ for association, even though they did not reach the required significance level. For example, rs9880192, located in the intergenic region between *c3orf27* and rs1991866, an intergenic variant at 8q24.21, shows p values $< 10^{-6}$ for monocyte level and $< 10^{-3}$ for MLR.

eQTL Results

Among the significant GWA loci for MLR and blood cell counts, there were a number of associations between the SNPs of interest and nearby gene expression (Table 6). However, the SNPs identified in our GWAS have low LD ($r^2 < 0.8$) with the top SNPs associated with gene expression, which suggest the GWAS SNPs are not part of the functional eQTL locus. Furthermore, no eQTLs with trans-effects were identified. In conclusion, we did not detect any cis- or trans-effects for the SNPs of interest.

GCTA and LD Score Regression

The results of the GCTA are shown in Table 7. From the GCTA we found a narrow-sense heritability of 43.3% for MLR, 54.1% for monocyte count, and 51.7% for lymphocyte count. The significant SNPs obtained in the GWAS for MLR explained 0.6% of the variance in MLR, and the significant SNPs obtained in the GWAS for monocyte count explained 4.4% of the variance in monocyte count. All known loci from published literature together explained 1.3% of MLR variance, 2.4 % of monocyte count variance and 0.3% of lymphocyte count variance. In LD score regression, all λ values were larger than the LD score regression intercept, and intercepts were close to 1, indicating that the inflation of the p value distribution from the GWAS results is caused by polygenetic effects, rather than population stratification. The SNP heritability of MLR, monocyte count, and lymphocyte count when applying LD score regression was 13%, 17%, and 19%, respectively (see Table 8).

In addition, we observed positive phenotypic correlations between MLR and monocyte count ($r = 0.550$, $p < .0001$) and between monocyte count and lymphocyte count ($r = 0.386$, $p < .0001$), and a negative phenotypic correlation between MLR and lymphocyte count ($r = -0.494$, $p < .0001$). However, despite the presence of phenotypic

TABLE 7

Narrow Sense Heritability (Standard Error) and the Proportion of Genetic Variance Explained by Known Significant SNPs (Standard Error) According to GCTA Analyses for Monocyte Count Ratio and its Subcomponents

Phenotype	Narrow-sense heritability (SE)	p	Proportion of genetic variance explained by significant SNPs (SE)	p	Proportion of genetic variance explained by known loci (SE)	p
MLR	0.43358 (0.025)	3.0E-8	0.005828 (0.008)	9.6E-9	0.01329 (0.006)	4.1E-13
Monocyte	0.5408 (0.022)	2.0E-12	0.043714 (0.024)	8.5E-5	0.02342 (0.009)	5.9E-11
Lymphocyte	0.51737 (0.023)	4.9E-8	NA	NA	0.00274 (0.002)	0.014

Note: p values in bold type indicate significant estimates.

TABLE 8

LD Score Regression Results for MLR, Monocyte Count and Lymphocyte Count

	MLR	Monocyte	Lymphocyte
Median of signed_ sumstats	−9.23E-05	9.03E-06	−6.58E-05
Mean of χ^2	1.007	1.011	1.021
λ GC	1.002	1.017	1.018
h^2 (se)	0.1302 (0.0733)	0.1702 (0.0854)	0.1912 (0.0895)
Intercept (se)	0.9915 (0.0063)	0.9912 (0.0073)	1.0011 (0.0069)

Note: Estimates for heritability and intercept with SE in brackets. All estimates are significant ($p > 0.05$).

associations, no significant genetic correlations were detected between any pair of variables.

Discussion

We presented a detailed examination of the causes of variance in MLR in the general population. Health status was, as expected, an important determinant of MLR level: individuals with a compromised immune system, our so-called unhealthy group, had on average a higher MLR than the healthy participants. In the healthy population, smoking, sex, and age and their interactions were important determinants of variation in MLR and its subcomponents. Smoking was associated with a higher MLR and this is in line with the higher MLR in the individuals with a compromised immune system in our study and the higher MLR seen in cancer patients. Genetic factors also were shown to play a role in MLR variation in the general population. Heritability for MLR was estimated at 40% and MLR level was associated with a locus near *ITGA4*. Earlier studies have shown this locus to be associated with monocyte levels. Heritability estimates were higher for MLR subcomponents (58% for both lymphocyte and monocyte count) and, in contrast to MLR, there was evidence for non-additive effects. Monocyte level was also associated with *ITGA4*, and four more genes were related to monocyte level in our analyses, replicating findings in previous GWA studies. For lymphocyte level, no significant genetic variants emerged.

From these results, it is clear that the genetic variants associated with blood cell counts may also influence their balance, as reflected in their ratios. In addition to the genetic variants in *ITGA4* genes, which were significantly associated with both MLR and monocyte count, there were

a number of loci that were significantly associated with monocyte count that may also affected MLR level: the loci near *LPARI*, *IRF8*, and *ITPR3* were marginally significant associated with MLR level. Also, a locus near *c3orf27* was marginally significantly associated with both MLR and monocyte count. We did not see evidence for pleiotropic genetic variants associated with MLR and lymphocyte count. To understand the role of the genetic variants in MLR variation, we investigated what is known about the role the identified genetic variants play in regulating gene expression. However, we did not find any evidence for cis-effects or trans-effects of these genetic variants.

Among three phenotypes, the narrow-sense heritability h^2 of lymphocyte (35.3% in the healthy population) was the lowest, but its SNP heritability was the highest (19.12% from the LD score regression). These results suggest more common autosomal SNPs may be associated with lymphocyte count. The LD score regression results showed that polygenic effects, rather than confounding factors, explain both ratio and count variance. Although there are significant overall correlations and an overlap in associated genetic variants has been detected between MLR and monocyte count, no significant genetic correlations were detected among variables, suggesting that the polygenic effects are too small to be detected with the current sample size.

Overall, this series of studies provided more insight into the causes of variation in MLR within the general population. While the genetic pathways as well as non-genetic causes of variance still need more clarification, it is clear that these factors need to be taken into consideration when studying the relationship between MLR and disease development.

Acknowledgments

BD Lin is supported by a PhD grant (201206180099) from the China Scholarship Council. This study was supported by multiple grants from the Netherlands Organization for Scientific Research (NWO: 016-115-035, 463-06-001, and 451-04-034), ZonMW (31160008 and 911-09-032); Institute for Health and Care Research (EMGO+) and Neuroscience Campus Amsterdam (NCA); Biomolecular Resources Research Infrastructure (BBMRI-NL, 184.021.007), European Research Council (ERC-230374); Biobanking and Biomolecular Resources

Research Infrastructure (BBMRINL: 184.021.007). Genotyping was made possible by grants from NWO/SPI 56-464-14192, Genetic Association Information Network (GAIN) of the Foundation for the National Institutes of Health, Rutgers University Cell and DNA Repository (NIMH U24 MH068457-06), the Avera Institute, Sioux Falls (USA) and the National Institutes of Health (NIH R01 HD042157-01A1, MH081802, Grand Opportunity grants 1RC2 MH089951 and 1RC2 MH089995).

References

- Abdellaoui, A., Hottenga, J. J., de Knijff, P., Nivard, M. G., Xiao, X., & Scheet, P. (2013). Population structure, migration, and diversifying selection in the Netherlands. *European Journal of Human Genetics*, *21*, 1277–1285.
- Al-Sufyani, A. A., & Mahassni, S. H. (2011). Obesity and immune cells in Saudi females. *Innate Immunity*, *17*, 439–450.
- Alvarez-Navarro, C., Martin-Esteban, A., Barnea, E., Admon, A., & de Castro, J. A. L. (2015). Endoplasmic reticulum aminopeptidase 1 (ERAP1) polymorphism relevant to inflammatory disease shapes the peptidome of the bird-shot chorioretinopathy-associated HLA-A*29:02 antigen. *Molecular & Cellular Proteomics*, *14*, 1770–1780.
- Azab, B., Camacho-Rivera, M., & Taioli, E. (2014). Average values and racial differences of neutrophil lymphocyte ratio among a nationally representative sample of United States subjects. *PLoS One*, *9*, e112361.
- Boker, S., Neale, M., Maes, H., Wilde, M., Spiegel, M., & Brick, T. (2011). OpenMx: An open source extended structural equation modeling framework. *Psychometrika*, *76*, 306–317.
- Boomsma, D. I., Willemsen, G., Sullivan, P. F., Heutink, P., Meijer, P., & Sondervan, D. (2008). Genome-wide association of major depression: Description of samples for the GAIN major depressive disorder study: NTR and NESDA biobank projects. *European Journal of Human Genetics*, *16*, 335–342.
- Bulik-Sullivan, B. K., Loh, P. R., Finucane, H. K., Ripke, S., Yang, J., & Patterson, N. (2015). LD score regression distinguishes confounding from polygenicity in genome-wide association studies. *Nature Genetics*, *47*, 291–295.
- Bulik-Sullivan, B., Finucane, H. K., Anttila, V., Gusev, A., Day, F. R., & Loh, P. R. (2015). An atlas of genetic correlations across human diseases and traits. *Nature Genetics*, *47*, 1236–1241.
- Crosslin, D. R., McDavid, A., Weston, N., Zheng, X. W., Hart, E., & Andrade, M. (2013). Genetic variation associated with circulating monocyte count in the eMERGE Network. *Human Molecular Genetics*, *22*, 2119–2127.
- De Jager, P. L., Jia, X., Wang, J., de Bakker, P. I., Ottoboni, L., & Aggarwal, N. T. (2009). Meta-analysis of genome scans and replication identify CD6, IRF8 and TNFRSF1A as new multiple sclerosis susceptibility loci. *Nature Genetics*, *41*, 776–782.
- Evans, D. M., Frazer, I. H., & Martin, N. G. (1999). Genetic and environmental causes of variation in basal levels of blood cells. *Twin Research*, *2*, 250–257.
- Farragher, T. M., Goodson, N. J., Naseem, H., Silman, A. J., Thomson, W., & Symmons, D. (2008). Association of the HLA-DRB1 gene with premature death, particularly from cardiovascular disease, in patients with rheumatoid arthritis and inflammatory polyarthritis. *Arthritis and Rheumatism*, *58*, 359–369.
- Ferreira, M. A., Hottenga, J. J., Warrington, N. M., Medland, S. E., Willemsen, G., & Lawrence, R. W. (2009). Sequence variants in three loci influence monocyte counts and erythrocyte volume. *American Journal of Human Genetics*, *85*, 745–749.
- Finucane, H. K., Bulik-Sullivan, B., Gusev, A., Trynka, G., Reshef, Y., & Loh, P. R. (2015). Partitioning heritability by functional annotation using genome-wide association summary statistics. *Nature Genetics*, *47*, 1228–1235.
- Hall, M. A., Ahmadi, K. R., Norman, P., Snieder, H., MacGregor, A. J., & Vaughan, R. W. (2000). Genetic influence on peripheral blood T lymphocyte levels. *Genes & Immunity*, *1*, 423–427.
- Hanifin, J. M., & Cline, M. J. (1970). Human monocytes and macrophages. Interaction with antigen and lymphocytes. *Journal of Cell Biology*, *46*, 97–105.
- Howie, B., Fuchsberger, C., Stephens, M., Marchini, J., & Abecasis, G. R. (2012). Fast and accurate genotype imputation in genome-wide association studies through pre-phasing. *Nat Genet*, *44*(8), 955–959.
- Iqbal, S., Umbreen, A., & Zaidi, S. B. H. (2014). Monocyte lymphocyte ratio as a possible prognostic marker in antituberculous therapy. *JRMC*, *18*(2), 178–181.
- Jansen, R., Hottenga, J. J., Nivard, M., Abdellaoui, A., Laport, B., & de Geus, E. J. (in press). Conditional eQTL analysis reveals allelic heterogeneity of gene expression. *Human Molecular Genetics*.
- Kamatani, Y., Matsuda, K., Okada, Y., Kubo, M., Hosono, N., & Daigo, Y. (2010). Genome-wide association study of hematological and biochemical traits in a Japanese population. *Nature Genetics*, *42*, 210–215.
- Keller, M. F., Reiner, A. P., Okada, Y., van Rooij, F. J., Johnson, A. D., & Chen, M. H. (2014). Trans-ethnic meta-analysis of white blood cell phenotypes. *Human Molecular Genetics*, *23*, 6944–6960.
- Kurotaki, D., Osato, N., Nishiyama, A., Yamamoto, M., Ban, T., & Sato, H. (2013). Essential role of the IRF8-KLF4 transcription factor cascade in murine monocyte differentiation. *Blood*, *121*, 1839–1849.
- Li, J., Chen, Q. Y., Luo, X. H., Hong, J., Pan, K. Y., & Lin, X. H. (2015). Neutrophil-to-lymphocyte ratio positively correlates to age in healthy population. *Journal of Clinical Laboratory Analysis*, *29*, 437–443.
- Li, Y., Willer, C. J., Ding, J., Scheet, P., & Abecasis, G. R. (2010). MaCH: Using sequence and genotype data to estimate haplotypes and unobserved genotypes. *Genetic Epidemiology*, *34*, 816–834.
- Lin, B. D., Hottenga, J. J., Abdellaoui, A., Dolan, C. V., de Geus, E. J., & Kluff, C. (2016). Causes of variation in the neutrophil-lymphocyte and platelet-lymphocyte ratios: A twin-family study. *Biomarkers In Medicine*, *10*, 1061–1072.

- Lin, B. D., Montoro, E., Bell, J., Boomsma, D. I., de Geus, E. J., & Jansen, R. (2016). *SNP heritability and effects of genetic variants for neutrophil-to-lymphocyte and platelet-to-lymphocyte ratio*. Manuscript submitted for publication.
- Liu, E. Y., Li, M., Wang, W., & Li, Y. (2013). MaCH-admix: Genotype imputation for admixed populations. *Genetic Epidemiology*, *37*, 25–37.
- Maugeri, N., Powell, J. E., Hoen, P. A., de Geus, E. J., Willemsen, G., & Kattenberg, M. (2011). LPAR1 and ITGA4 regulate peripheral blood monocyte counts. *Human Mutation*, *32*, 873–876.
- Nalls, M. A., Couper, D. J., Tanaka, T., van Rooij, F. J., Chen, M. H., & Smith, A. V. (2011). Multiple loci are associated with white blood cell phenotypes. *Plos Genetics*, *7*, e1002113.
- Nishijima, T. F., Muss, H. B., Shachar, S. S., Tamura, K., & Takamatsu, Y. (2015). Prognostic value of lymphocyte-to-monocyte ratio in patients with solid tumors: A systematic review and meta-analysis. *Cancer Treatment Reviews*, *41*, 971–978.
- Oishi, T., Iida, A., Otsubo, S., Kamatani, Y., Usami, M., & Takei, T. (2008). A functional SNP in the NKX2.5-binding site of ITPR3 promoter is associated with susceptibility to systemic lupus erythematosus in Japanese population. *Journal of Human Genetics*, *53*, 151–162.
- Okada, Y., Hirota, T., Kamatani, Y., Takahashi, A., Ohmiya, H., & Kumasaka, N. (2011). Identification of nine novel loci associated with white blood cell subtypes in a Japanese population. *PLOS Genetics*, *7*(6), e1002067.
- Pe'er, I., Yelensky, R., Altshuler, D., & Daly, M. J. (2008). Estimation of the multiple testing burden for genomewide association studies of nearly all common variants. *Genetic Epidemiology*, *32*, 381–385. doi:10.1002/gepi.20303.
- Perez-de-Heredia, F., Gomez-Martinez, S., Diaz, L. E., Veses, A. M., Nova, E., & Warnberg, J. (2015). Influence of sex, age, pubertal maturation and body mass index on circulating white blood cell counts in healthy European adolescents—the HELENA study. *European Journal of Pediatrics*, *174*, 999–1014.
- R Core Team. (2014). *R: a language and environment for statistical computing*. Vienna, Austria: R Foundation for Statistical Computing.
- Reiner, A. P., Lettre, G., Nalls, M. A., Ganesh, S. K., Mathias, R., & Austin, M. A. (2011). Genome-wide association study of white blood cell count in 16,388 African Americans: The Continental Origins and Genetic Epidemiology Network (COGENT). *PLOS Genetics*, *7*, e1002108.
- Schwartz, J., & Weiss, S. T. (1994). Cigarette smoking and peripheral blood leukocyte differentials. *Annals of Epidemiology*, *4*, 236–242.
- Shabalin, A. A. (2012). Matrix eQTL: Ultra fast eQTL analysis via large matrix operations. *Bioinformatics*, *28*, 1353–1358.
- Sirota, M., Willemsen, G., Sundar, P., Pitts, S. J., Potluri, S., & Prifti, E. (2015). Effect of genome and environment on metabolic and inflammatory profiles. *PLoS One*, *10*, e0120898.
- Stata Corp. (2013). *Stata statistical software: Release 13*. College Station, TX: StataCorp LP.
- Tenorio, T. R., Farah, B. Q., Ritti-Dias, R. M., Botero, J. P., Brito, D. C., & Moura, P. M. (2014). Relation between leukocyte count, adiposity, and cardiorespiratory fitness in pubertal adolescents. *Einstein*, *12*, 420–424.
- Terry, R. L., Deffrasnes, C., Getts, D. R., Minten, C., van Vreden, C., & Ashhurst, T. M. (2015). Defective inflammatory monocyte development in IRF8-deficient mice abrogates migration to the west Nile virus-infected brain. *Journal of Innate Immunity*, *7*, 102–112.
- Thomas, D., Gazouli, M., Karantanos, T., Rigoglou, S., Karamanolis, G., & Bramis, K. (2014). Association of rs1568885, rs1813443 and rs4411591 polymorphisms with anti-TNF medication response in Greek patients with Crohn's disease. *World Journal of Gastroenterology*, *20*, 3609–3614.
- Tucker, G., Loh, P. R., MacLeod, I. M., Hayes, B. J., Goddard, M. E., & Berger, B. (2015). Two-variance component model improves genetic prediction in family datasets. *American Journal of Human Genetics*, *97*, 677–690.
- Willemsen, G., de Geus, E. J., Bartels, M., van Beijsterveldt, C. E., Brooks, A. I., & Estourgie-van Burk, G. F. (2010). The netherlands twin register biobank: A resource for genetic epidemiological studies. *Twin Research and Human Genetics*, *13*, 231–245.
- Willemsen, G., Vink, J. M., Abdellaoui, A., den Braber, A., van Beek, J. H., & Draisma, H. H. (2013). The adult netherlands twin register: Twenty-five years of survey and biological data collection. *Twin Research and Human Genetics*, *16*(1), 271–281.
- Wright, F. A., Sullivan, P. F., Brooks, A. I., Zou, F., Sun, W., & Xia, K. (2014). Heritability and genomics of gene expression in peripheral blood. *Nature Genetics*, *46*, 430–437.
- Yanez, A., Ng, M. Y., Hassanzadeh-Kiabi, N., & Goodridge, H. S. (2015). IRF8 acts in lineage-committed rather than oligopotent progenitors to control neutrophil vs monocyte production. *Blood*, *125*, 1452–1459.
- Yang, J., Lee, S. H., Goddard, M. E., & Visscher, P. M. (2011). GCTA: A tool for genome-wide complex trait analysis. *American Journal of Human Genetics*, *88*, 76–82.
- Zaitlen, N., Kraft, P., Patterson, N., Pasaniuc, B., Bhatia, G., Pollack, S., & Price, A. L. (2013). Using extended genealogy to estimate components of heritability for 23 quantitative and dichotomous traits. *PLOS Genetics*, *9*, e1003520.
- Zaldivar, F., McMurray, R. G., Nemet, D., Galassetti, P., Mills, P. J., & Cooper, D. M. (2006). Body fat and circulating leukocytes in children. *International Journal of Obesity*, *30*, 906–911.